Expectation and Variance

Analysis of Common Distributions

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Expectation of Bernoulli Trials

Probability Mass Function:

$$f_x(x) = p^x (1-p)^{1-x} \text{ for } x \in \{0, 1\}$$

Expectation of X:

$$E(X) = \sum_{x=0}^{x=1} x f_x(x) = \sum_{x=0}^{x=1} x [p^x (1-p)^{1-x}]$$

$$= 0. p^0 (1-p)^1 + 1. p^1 (1-p)^0$$

$$= 0 + p$$

$$= p$$

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Variance of Bernoulli Trials

$$E(X^{2}) = \sum_{x=0}^{x=1} x^{2} f_{x}(x) = \sum_{x=0}^{x=1} x^{2} [p^{x} (1-p)^{1-x}]$$

$$= 0. p^{0} (1-p)^{1} + 1. p^{1} (1-p)^{0}$$

$$= 0 + p$$

$$= p$$

Variance of Bernoulli Trials

$$Var(X) = E(X^{2}) - E(X)^{2}$$
$$= p - p^{2}$$
$$= p(1 - p)$$

Expectation of Binomial Distribution

Probability Mass Function:

 $f_x(x) = \binom{n}{x} \cdot p^x (1-p)^{n-x}$ for $x \in W$ where W are all the natural numbers.

$$E(X) = \sum_{x=0}^{x=n} x f_x(x)$$
$$= \sum_{x=0}^{x=n} x \binom{n}{x} \cdot p^x (1-p)^{n-x}$$

Expectation of X:

But this is too difficult. Therefore recall that the Binomial Distribution is the sum of Bernoulli trials. Therefore, $X = \sum_{i=1}^{n} Y_i$ where X is the Binomial random variable and Y_i is the Bernoulli random variable.

$$E(X) = E\left[\sum_{i=1}^{n} Y_{i}\right]$$

$$= \sum_{i=1}^{n} E[Y_{i}]$$

$$= \sum_{i=1}^{n} p$$

$$= np$$

Variance of Binomial Distribution

We will calculate the variance of a Binomial Distribution in the same way that we calculated the expectation.

$$Var(X) = Var \left[\sum_{i=1}^{n} Y_i \right]$$
$$= \sum_{i=1}^{n} Var[Y_i]$$
$$= \sum_{i=1}^{n} p(1-p)$$
$$= np(1-p)$$

Expectation of Geometric Distribution

Probability Mass Function:

 $f_x(x) = p(1-p)^x$ for $x \in N$ where N are all the natural numbers.

$$E(X) = \sum_{x=0}^{\infty} x \cdot f_x(x)$$
$$= \sum_{x=0}^{\infty} x \cdot p(1-p)^{x-1}$$
$$= p \cdot \sum_{x=0}^{\infty} x \cdot (1-p)^{x-1}$$

Expectation of Geometric Distribution

Notice that $\sum_{n=0}^{\infty} y^n = \frac{1}{1-y}$.

This is the Taylor Series Expansion of $y \in (0,1)$. Differentiating both sides with $\frac{d}{dv}$:

$$\sum_{n=0}^{\infty} n \cdot y^{n-1} = \frac{1}{(1-y)^2}$$

. Now put n = x and y = (1 - p), so we get:

$$\sum_{x=0}^{\infty} x \cdot (1-p)^{x-1} = \frac{1}{(1-(1-p))^2} = \frac{1}{p^2}$$

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Expectation of Geometric Distribution

Expectation of X: We put the value found in the previous slide on this slide.

$$E(X) = p \cdot \sum_{x=0}^{\infty} x \cdot (1 - p)^{x-1}$$
$$= p \cdot \frac{1}{p^2}$$
$$= \frac{1}{p}$$

Expectation of Uniform Distribution

Probability Density Function:

$$f_x(x) = \frac{1}{h-a}$$
 for $a, b \in R$ where R are all the real numbers.

$$E(X) = \int x \cdot f_x(x) dx$$
$$= \int_a^b x \cdot \frac{1}{b-a} dx$$
$$= \frac{1}{(b-a)} \int_a^b x dx$$

$$E(X) = \frac{1}{(b-a)} \left[\frac{1}{2} \cdot x^2 \Big|_a^b \right]$$

$$= \frac{1}{(b-a)} \left[\frac{b^2 - a^2}{2} \right]$$

$$= \frac{1}{(b-a)} \cdot \frac{(b+a)(b-a)}{2}$$

$$= \frac{(b+a)}{2}$$

Expectation of Exponential Distribution

Probability Density Function:

 $f_x(x) = \lambda \cdot \exp^{-\lambda x}$ for $x \in R_+$ where R_+ are all the positive real numbers.

$$E(X) = \int x \cdot f_x(x) dx$$
$$= \int_0^\infty x \cdot \lambda \cdot \exp^{-\lambda x} dx$$

Expectation of Exponential Distribution

We will use by parts integration to solve this problem.

$$\int v \, du = uv - \int u \, dv$$

$$v = x$$

$$dv = dx$$

$$u = \exp^{-\lambda x}$$

$$du = -\lambda \cdot \exp^{-\lambda x} dx$$
(1)

Expectation of Exponential Distribution

$$E(X) = \int_0^\infty x \cdot \lambda \cdot \exp^{-\lambda x} dx$$

$$= -\left[x \cdot \exp^{-\lambda x}\Big|_0^\infty\right] + \int_0^\infty \exp^{-\lambda x} dx$$

$$= -\left[\infty \cdot (\exp^{-\lambda \cdot \infty}) - 0 \cdot (\exp^{-\lambda \cdot 0})\right] + \left[\frac{-1}{\lambda} \exp^{-\lambda x}\Big|_0^\infty\right]$$

$$= \left[\left(\frac{-1}{\lambda} \exp^{-\lambda \cdot \infty}\right) - \left(\frac{-1}{\lambda} \exp^{-\lambda \cdot 0}\right)\right]$$

$$= \frac{1}{\lambda}$$

Probability Density Function:

$$f_{x}(x)=rac{1}{\sqrt{2\pi}}\cdot \exp^{-rac{1}{2}x^{2}}$$
 for $x\in R$ where R are all the real numbers.

$$E(X) = \int x \cdot f_x(x) dx$$
$$= \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi}} \exp^{-\frac{1}{2}x^2} dx$$

We will use substitution to solve this problem.

$$u = x^{2}$$

$$du = 2x dx$$

$$dx = \frac{du}{2x}$$
(2)

$$E(X) = \int_{-\infty}^{\infty} x \cdot \frac{1}{\sqrt{2\pi}} \exp^{-\frac{1}{2}x^2} dx$$

$$= \frac{1}{\sqrt{2\pi}} \cdot \int_{-\infty}^{\infty} x \cdot \exp^{-\frac{1}{2}x^2} dx$$

$$= \frac{1}{\sqrt{2\pi}} \cdot \int_{-\infty}^{\infty} x \cdot \exp^{-\frac{1}{2}u} \frac{du}{2x}$$

$$= \frac{1}{\sqrt{2\pi}} \cdot \int_{-\infty}^{\infty} \frac{1}{2} \cdot \exp^{-\frac{1}{2}u} du$$

$$= \frac{1}{\sqrt{2\pi}} \left[\frac{1}{2} \cdot -2 \cdot \exp^{-\frac{1}{2}u} \right]_{-\infty}^{\infty}$$

$$E(X) = \frac{1}{\sqrt{2\pi}} \left[-\exp^{-\frac{1}{2}x^2} \Big|_{-\infty}^{\infty} \right]$$

$$= \frac{1}{\sqrt{2\pi}} \left[-\exp^{-\frac{1}{2}\infty} - \left(-\exp^{-\frac{1}{2}\infty} \right) \right]$$

$$= \left[-\exp^{-\frac{1}{2}\infty} + \exp^{-\frac{1}{2}\infty} \right]$$

$$= 0$$

Expectation of Normal Distribution

The Normal random variable X is expressed as $X = \mu + \sigma Z$ where Z is the Standard Normal Distribution.

$$E(X) = E(\mu + \sigma Z)$$

$$= \mu + \sigma \cdot E(Z)$$

$$= \mu + \sigma \cdot 0$$

$$= \mu$$

End

We will not discuss the variance of every distribution because the calculations are too cumbersome, however you can practice on your own calculation using $Var(X) = E(X^2) - E(X)^2$ for each case.